# Machine Learning 

## -Linear Regression-

SCH Univ.
Dept. of AI and Bigdata
Kim JinSeong

## Contents

1. Chapter I

- Definition of Linear Regression

2. Chapter $\Pi$

- Parameter Estimation

3. Chapter III

- Parameter Inference

4. Chapter IV

- Coefficient of Determination
- ANOVA


## Chapter I

-Definition of Linear Regression-

## Types of Machine Learning



## Types of Regression Model



## Types of Regression Model



## Definition of Linear Regression Model

Linear Regression Model: Model that expresses Y(output variable) as a linear combination of X(input variable)

* Linear combination : Combine variables by adding/subtracting (constant multiplication)
ex) $Y=B_{0}+B_{1} X_{1}+B_{2} X_{2}+\cdots+B_{p} X_{p}$


Purpose 1. Explain the relationship between $X$ variables and $Y$ variables
2. Predict future $Y$ (output variables)

## Definition of Linear Regression Model



$$
\begin{array}{r}
Y=\frac{\text { can be explained by } X(f(x))}{}+\frac{\text { can't be explained by } X(\varepsilon)}{\varepsilon=\text { random error }}
\end{array}
$$

## Assumption of Linear Regression Model

- Assumption of random error
$\Rightarrow \varepsilon_{\mathrm{i}} \sim \mathrm{N}\left(0, \sigma^{2}\right) \quad \mathrm{i}=1,2,3, \ldots, \mathrm{n}$
$\varepsilon_{\mathrm{i}}$ conforms to a normal distribution $\rightarrow \mathrm{E}\left(\varepsilon_{\mathrm{i}}\right)=0, \mathrm{~V}\left(\varepsilon_{\mathrm{i}}\right)=\sigma^{2}$ for all i

In $\mathrm{Y}=B_{0}+B_{1} \mathrm{X}+\varepsilon, \varepsilon$ Follows probability distribution(normal distribution) So, Y also follows any probability distribution

$$
\begin{array}{ll}
\text { 1. } \mathrm{E}\left(Y_{\mathrm{i}}\right)=\mathrm{E}\left(B_{0}+B_{1} \mathrm{X}_{\mathrm{i}}\right)+\mathrm{E}(\varepsilon)=B_{0}+B_{1} \mathrm{X}_{\mathrm{i}} & \text { 2. } \mathrm{V}\left(Y_{\mathrm{i}}\right)=\mathrm{V}\left(B_{0}+B_{1} \mathrm{X}_{\mathrm{i}}\right)+\mathrm{V}(\varepsilon)=\sigma^{2} \\
& \text { B } B_{0}+B_{1} \mathrm{X}_{\mathrm{i}} \text { is constant } \rightarrow \mathrm{E}\left(B_{0}+B_{1} \mathrm{X}_{\mathrm{i}}\right)=B_{0}+B_{1} \mathrm{X}_{\mathrm{i}} \\
\mathrm{E}\left(\varepsilon_{\mathrm{i}}\right)=0 & {\left[\begin{array}{l}
B_{0}+B_{1} \mathrm{X}_{\mathrm{i}} \text { is constant } \rightarrow \mathrm{V}\left(B_{0}+B_{1} \mathrm{X}_{\mathrm{i}}\right)=0 \\
\mathrm{~V}\left(\varepsilon_{\mathrm{i}}\right)=\sigma^{2}
\end{array}\right.}
\end{array}
$$

$$
\text { i.e., } Y_{i} \sim N\left(B_{0}+B_{1} X_{i}, \sigma^{2}\right) i=1,2, \cdots, n
$$

## Assumption of Linear Regression Model



$$
\text { i.e., } Y_{i} \sim N\left(B_{0}+B_{1} X_{i}, \sigma^{2}\right) i=1,2, \cdots, n
$$

## Linear Regression Model



## View Point.

Find a linear regression line that describes the relationship
between the input variable $(\mathrm{X})$ and the mean of output variable $(\mathrm{Y})$
i.e., Find Parameter $\left(B_{0}, B_{1}, \ldots, B_{p}\right)$ using the function of data

## Linear Regression Model



View Point.
Find a linear regression line that describes the relationship between the input variable $(\mathrm{X})$ and the mean of output variable $(\mathrm{Y})$

Linear Regression Model


## Linear Regression Model



## Linear Regression Model



Find Best Parameter $\left(B_{0}, B_{1}, ., B_{p}\right)$ using data

## Linear Regression Model



Find Best Parameter $\left(B_{0}, B_{1}, \ldots, B_{p}\right)$ using data
How to find good parameter?

## Chapter $\Pi$

- Parameter Estimation -


## Parameter Estimation

Question. Let's compare with red and blue. Which one is correct prediction line?


## Parameter Estimation

Question. Let's compare with red and blue. Which one is correct prediction line?


## Parameter Estimation

Question. Let's compare with red and blue. Which one is correct prediction line?


Answer. Red is a better regression line than blue

## Parameter Estimation



$$
\begin{aligned}
& \mathrm{d}_{1}+\mathrm{d}_{2}+\cdots+\mathrm{d}_{\mathrm{n}}=0 \\
& d_{1}^{2}+d_{2}^{2}+\cdots+d_{\mathrm{n}}^{2} \geq 0 \\
& \mathrm{~d}_{1}=\mathrm{Y}_{1}-E\left(\mathrm{Y}_{1}\right) \\
& =\mathrm{Y}_{1}-\left(\mathrm{B}_{0}+\mathrm{B}_{1} \mathrm{X}_{1}\right) \\
& \sum_{i=1}^{n} d_{\mathrm{i}}^{2}=\sum_{i=1}^{n}\left\{\mathrm{Y}_{\mathrm{i}}-\left(\mathrm{B}_{0}+\mathrm{B}_{1} \mathrm{X}_{\mathrm{i}}\right)\right\}^{2} \longleftarrow \text { Cost Function }
\end{aligned}
$$

i.e., Finding the smallest Cost function is finding the best parameters !!!

$$
\min _{\mathrm{B}_{0}, \mathrm{~B}_{1}} \sum_{i=1}^{n}\left\{\mathrm{Y}_{\mathrm{i}}-\left(\mathrm{B}_{0}+\mathrm{B}_{1} \mathrm{X}_{\mathrm{i}}\right)\right\}^{2}
$$

## Parameter Estimation



$$
\begin{aligned}
& \mathrm{d}_{1}+\mathrm{d}_{2}+\cdots+\mathrm{d}_{\mathrm{n}}=0 \\
& d_{1}^{2}+d_{2}^{2}+\cdots+d_{\mathrm{n}}^{2} \geq 0 \\
& \mathrm{~d}_{1}=\mathrm{Y}_{1}-E\left(\mathrm{Y}_{1}\right) \\
& =Y_{1}-\left(B_{0}+B_{1} X_{1}\right) \\
& \sum_{i=1}^{n} d_{\mathrm{i}}^{2}=\sum_{i=1}^{n}\left\{\mathrm{Y}_{\mathrm{i}}-\left(\mathrm{B}_{0}+\mathrm{B}_{1} \mathrm{X}_{\mathrm{i}}\right)\right\}^{2} \longleftarrow \text { Cost Function }
\end{aligned}
$$

i.e., Finding the smallest Cost function is finding the best parameters !!!

$$
\min _{\mathrm{B}_{0}, \mathrm{~B}_{1}} \sum_{i=1}^{n}\left\{\mathrm{Y}_{\mathrm{i}}-\left(\mathrm{B}_{0}+\mathrm{B}_{1} \mathrm{X}_{\mathrm{i}}\right)\right\}^{2}
$$

## Parameter Estimation

In linear regression, Cost Function is always convex = globally optional solution exists


Convex Function


Local Optimal Solution
i.e., The way that finds the smallest cost function(estimates best parameter) is

## Parameter Estimation

- Partial derivative based on Parameter $\left(\mathrm{B}_{1}, \mathrm{~B}_{0}\right)$
( $\mathrm{B}_{1}$ : gradient, , $\mathrm{B}_{0}: \mathrm{y}$-intercept)
Cost Function: $\quad \sum_{i=1}^{n}\left\{\mathrm{Y}_{\mathrm{i}}-\left(\mathrm{B}_{0}+\mathrm{B}_{1} \mathrm{X}_{\mathrm{i}}\right)\right\}^{2}$
$\left[\begin{array}{ll}\mathrm{B}_{0} \text { partial derivative } & \rightarrow \frac{\partial \mathrm{C}\left(\mathrm{B}_{0}, \mathrm{~B}_{1}\right)}{\partial \mathrm{B}_{0}}=-2 \sum_{i=1}^{\mathrm{n}} \mathrm{Y}_{\mathrm{i}}-\left(\mathrm{B}_{0}+\mathrm{B}_{1} \mathrm{X}_{\mathrm{i}}\right)=0 \\ \mathrm{~B}_{1} \text { partial derivative } & \rightarrow \frac{\partial \mathrm{C}\left(\mathrm{B}_{0}, \mathrm{~B}_{1}\right)}{\partial \mathrm{B}_{1}}=-2 \sum_{\mathrm{i}=1}^{\mathrm{n}} \mathrm{Y}_{\mathrm{i}}-\left(\mathrm{B}_{0}+\mathrm{B}_{1} \mathrm{X}_{\mathrm{i}}\right) \mathrm{X}_{\mathrm{i}}=0\end{array}\right.$

The result of partial derivative

$$
\left[\begin{array}{l}
\hat{B}_{0}=\bar{Y}-\widehat{B}_{0} \bar{X} \\
\hat{B}_{1}=\frac{\sum_{\mathrm{i}=1}^{\mathrm{n}}\left(\mathrm{X}_{\mathrm{i}}-\overline{\mathrm{X}}\right)\left(\mathrm{Y}_{\mathrm{i}}-\overline{\mathrm{Y}}\right)}{\sum_{\mathrm{i}=1}\left(\mathrm{X}_{\mathrm{i}}-\overline{\mathrm{X}}\right)^{2}}
\end{array}\right.
$$

The linear regression function that has best parameter

$$
f(x)=\widehat{Y}=\widehat{B}_{0}+\widehat{B}_{1} X
$$

## Least Squares Estimation Algorithm

Goal. Find estimator of $\mathrm{B}_{0}$ and $\mathrm{B}_{1}$ (i.e., $\hat{B}_{0}$ and $\widehat{B}_{1}$ )
Step1. Cost Function(Squared the sum of the difference between the actual $y$ value and $y$ value on the regression line)

$$
\sum_{i=1}^{n}\left\{\mathrm{Y}_{\mathrm{i}}-\left(\mathrm{B}_{0}+\mathrm{B}_{1} \mathrm{X}_{\mathrm{i}}\right)\right\}^{2}
$$

Step2. Find $\mathrm{B}_{0}, \mathrm{~B}_{1}$ to minimize Cost Function

$$
\min _{\mathrm{B}_{0}, \mathrm{~B}_{1}} \sum_{i=1}^{n}\left\{\mathrm{Y}_{\mathrm{i}}-\left(\mathrm{B}_{0}+\mathrm{B}_{1} \mathrm{X}_{\mathrm{i}}\right)\right\}^{2}
$$

Step3. Find the point where the derivative(gradient) is 0

$$
\begin{gathered}
\frac{\partial \mathrm{C}\left(\mathrm{~B}_{0}, \mathrm{~B}_{1}\right)}{\partial \mathrm{B}_{0}}=-2 \sum_{\mathrm{i}=1}^{\mathrm{n}} \mathrm{Y}_{\mathrm{i}}-\left(\mathrm{B}_{0}+\mathrm{B}_{1} \mathrm{X}_{\mathrm{i}}\right)=0 \\
\frac{\partial \mathrm{C}\left(\mathrm{~B}_{0}, \mathrm{~B}_{1}\right)}{\partial \mathrm{B}_{1}}=-2 \sum_{\mathrm{i}=1}^{\mathrm{n}} \mathrm{Y}_{\mathrm{i}}-\left(\mathrm{B}_{0}+\mathrm{B}_{1} \mathrm{X}_{\mathrm{i}}\right) \mathrm{X}_{\mathrm{i}}=0
\end{gathered}
$$

Solutions. $\hat{B}_{0}=\bar{Y}-\hat{B}_{1} \bar{X}, \hat{B}_{1}=\frac{\sum_{\mathrm{i}=1}^{\mathrm{n}}\left(\mathrm{X}_{\mathrm{i}}-\overline{\mathrm{X}}\right)\left(\mathrm{Y}_{\mathrm{i}}-\overline{\mathrm{Y}}\right)}{\sum_{\mathrm{i}=1}^{\mathrm{n}}\left(\mathrm{X}_{\mathrm{i}}-\overline{\mathrm{X}}\right)^{2}}$

## Residual


$\left[\begin{array}{l}B_{0}, B_{1} \text { is not fixed value, just status of parameter } \\ \varepsilon\end{array}\right.$ $\varepsilon$ follows normal distribution

[ $\hat{B}_{0}, \widehat{B}_{1}$ is fixed value
e is error of fixed values (constant)

$$
\mathrm{e}(\text { residual })=\text { the value that } \varepsilon \text { (random error }) \text { is actually implemented }
$$

## Chapter III

- Parameter Inference -


## Parameter inference

- There are two ways of infer parameters

1. Estimator
2. Hypothesis test

## Estimator of parameter

- Estimators $\left(\widehat{B}_{0}, \widehat{B}_{1}\right)$ that calculated by using Least Squared Estimation Algorithm

$$
\hat{B}_{0}=\bar{Y}-\hat{B}_{1} \bar{X}, \quad \hat{B}_{1}=\frac{\sum_{\mathrm{i}=1}^{\mathrm{n}}\left(\mathrm{X}_{\mathrm{i}}-\overline{\mathrm{X}}\right)\left(\mathrm{Y}_{\mathrm{i}}-\overline{\mathrm{Y}}\right)}{\sum_{\mathrm{i}=1}^{\mathrm{n}}\left(\mathrm{X}_{\mathrm{i}}-\overline{\mathrm{X}}\right)^{2}}
$$

- Estimator: a function of the sample(data)
$\widehat{B}_{0}, \widehat{B}_{1}$
- Usage of Estimator: estimate unknown parameter $\left(\mathrm{B}_{0}, \mathrm{~B}_{1}\right)$
- Types of Estimator - Point Estimator
- Interval Estimator


## Point estimator of parameter

$$
Y_{\mathrm{i}}=B_{0}+B_{1} \mathrm{X}_{\mathrm{i}}+\varepsilon_{\mathrm{i}} \quad \varepsilon_{\mathrm{i}} \sim \mathrm{~N}\left(0, \sigma^{2}\right) \mathrm{i}=1,2, \cdots, \mathrm{n}
$$

1) Point Estimator of $B_{0}: \hat{B}_{0}=\bar{Y}-\widehat{B}_{1} \bar{X}$
2) Point Estimator of $B_{1}: \widehat{B}_{1}=\frac{\sum_{\mathrm{i}=1}^{\mathrm{n}}\left(\mathrm{X}_{\mathrm{i}}-\overline{\mathrm{X}}\right)\left(\mathrm{Y}_{\mathrm{i}}-\overline{\mathrm{Y}}\right)}{\sum_{\mathrm{i}=1}^{\mathrm{n}}\left(\mathrm{X}_{\mathrm{i}}-\overline{\mathrm{X}}\right)^{2}}$
3) Point Estimator of $\sigma^{2}: \widehat{\sigma}^{2}=\left(\frac{1}{n-2}\right) \sum_{i=1}^{n} e_{i}{ }^{2} \quad(n=$ number of samples, $e=$ residual $)$

Gauss-Markov Theorem: Least Square Estimator is the Best Linear Unbiased Estimator (BLUE)
BLUE : The BLUE is (1)unbiased estimator and (2)has the smallest average squared error(variance) compared to any unbiased estimators.
(1) unbiased estimator: $\mathrm{E}\left(\widehat{B}_{0}\right)=B_{0}, \mathrm{E}\left(\widehat{B}_{1}\right)=B_{1}$
(2) smallest variance estimator: $\mathrm{V}\left(\mathrm{a} \hat{B}_{0}\right) \leq \mathrm{V}(\mathrm{b} \hat{\theta}), \mathrm{V}\left(\mathrm{a} \hat{B}_{1}\right) \leq \mathrm{V}(\mathrm{b} \hat{\theta}) \quad \hat{\theta}$ : any other unbiased estimate

## Interval estimator of parameter

- Con(s) of interval estimate
$\rightarrow$ Estimate intervals to provide more flexible information
- Basic form that interval estimator of $\theta$ (parameter)

$$
\widehat{\theta}-C * \sigma(\widehat{\theta}) \leq \theta \leq \widehat{\theta}+C * \sigma(\widehat{\theta}) \quad \widehat{\theta}: \text { point estimator of } \theta
$$

i.e., have to know parameter(point estimator, constance, standard deviation)

1) Confidence interval for gradient $\left(B_{1}\right) \rightarrow(100(1-a) \%)$

$$
\Rightarrow \hat{B}_{1}-t_{\frac{a}{2}, n-2} s d\left(\hat{B}_{1}\right) \leq \hat{B}_{1} \leq \hat{B}_{1}+t_{\frac{a}{2}, n-2} s d\left(\hat{B}_{1}\right)
$$

1) $\widehat{B}_{1}=\frac{\sum_{\mathrm{i}=1}^{\mathrm{n}}\left(\mathrm{X}_{\mathrm{i}}-\overline{\mathrm{X}}\right)\left(\mathrm{Y}_{\mathrm{i}}-\overline{\mathrm{Y}}\right)}{\sum_{\mathrm{i}=1}^{\mathrm{n}}\left(\mathrm{X}_{\mathrm{i}}-\overline{\mathrm{X}}\right)^{2}}$ : point estimator of $B_{1}$
2) $t_{\frac{a}{2}, n-2}$ : The value of the $t$-distribution with a degree of freedom of $n-2$ under the significance level (1-a)
3) $\operatorname{sd}\left(\hat{B}_{1}\right)=\sqrt{\frac{\widehat{\sigma}^{2}}{\sum_{i=1}^{n}(X-\bar{X})^{2}}}$ : standard deviation of $\hat{B}_{1}$
4) Confidence interval for $y$-interval $\left(B_{0}\right)$
$\Rightarrow$ Same form as confidence interval for $B_{1}$

## Hypothesis test for gradient $\left(B_{1}\right)$

What is hypothesis test? Hypothesis and test for unknown parameters

- hypothesis test
$H_{0}: B_{1}=0$ vs $H_{1}: B_{1} \neq 0 \quad\left(H_{0}:\right.$ Null Hypothesis, $H_{1}$ : Alternative Hypothesis $)$
* If $B_{1}$ (gradient) $=0$, There is no relationship between X and Y

$$
t^{*}=\frac{\widehat{B}_{0}-0}{s d\left(\widehat{B}_{1}\right)} \leftarrow \text { test statistic for null hypothesis }
$$

( $\widehat{B}_{0}:$ made of data, $0:$ made of hypothesis, sd $\left(\widehat{B}_{1}\right):$ use for scaling)

Prove hypothesis test by one of the two methods

1) $I F\left|t^{*}\right|>t_{\frac{a}{2}, n-2} \rightarrow$ we reject $H_{0}$
2) p -value $=2 P\left(T>\left|t^{*}\right|\right)$ where $T \sim t(n-2)$

## Example (Regression analysis)

The regression equation $\longrightarrow \mathrm{Y}($ Appraised value $)=-29.6+0.0779 \mathrm{X}$ (Area)

| Predictor | Coef | SE Coef | T | P | S $=16.9065$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Constant | -29.59 | 10.66 | -2.78 | 0.016 |  |
| Area | 0.077939 | 0.004370 | 17.83 | 0.00 |  |

Q1. What are point estimates of the parameters?

$$
\Rightarrow \widehat{B}_{0}=-29.56, \widehat{B}_{1}=0.077939
$$

Q2. What is the standard deviation(standard error) of the parameter?

$$
\begin{aligned}
\Rightarrow s d\left(\hat{B}_{0}\right) & =\sqrt{\widehat{\sigma}^{2}\left[\frac{1}{n}+\frac{\bar{X}^{2}}{\sum_{i=1}^{n}\left(X_{i}-\bar{X}\right)^{2}}\right]}=10.66 \\
s d\left(\hat{B}_{1}\right) & =\sqrt{\frac{\hat{\sigma}^{2}}{\sum_{i=1}^{n}\left(X_{i}-\bar{X}\right)^{2}}}=0.004370
\end{aligned}
$$

## Example (Regression analysis)

The regression equation $\longrightarrow \mathrm{Y}($ Appraised value $)=-29.6+0.0779 \mathrm{X}$ (Area)

| Predictor | Coef | SE Coef | T | P | S $=16.9065$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Constant | -29.59 | 10.66 | -2.78 | 0.016 |  |
| Area | 0.077939 | 0.004370 | 17.83 | 0.00 |  |

Q3. What is the T in the above table?

$$
\begin{aligned}
& \Rightarrow H_{0}: B_{1}=0 \text { vs } H_{1}: B_{1} \neq 0 \\
& \mathrm{~T}=t^{*}=\frac{\widehat{B}_{0}-0}{s d\left(\hat{B}_{1}\right)}=\frac{0.077939-0}{0.004370}=17.83
\end{aligned}
$$

Q4. What is the $P$ in the above table?

$$
\begin{aligned}
& \Rightarrow>\mathrm{p}-\text { value }=2 P\left(T>\left|t^{*}\right|\right)=2 P(T>|17.83|) \text { where } T \sim t(13)(n=15 \rightarrow n-2=13)=0.00 \\
& \longrightarrow H_{0} \text { is rejected, } H_{1} \neq 0 \text { i.e., } \mathrm{X} \text { (Area) has significant effect on } \mathrm{Y} \text { (Appraised value) }
\end{aligned}
$$

Q5. What is the S in the above table?

$$
\Rightarrow \mathrm{S}=\widehat{\sigma}=\sqrt{\left(\frac{1}{n-2}\right) \sum_{i=1}^{n} e_{i}^{2}}=16.9065
$$

## Chapter IV

- Coefficient of Determination \& ANOVA-


## Coefficient of Determination: $\boldsymbol{R}^{\mathbf{2}}$



SSE (Sum of Square Error) $=\sum_{i=1}^{n}\left(Y_{i}-\widehat{Y}_{i}\right)^{2}$
$\operatorname{SSR}$ (Sum of Square Regression) $=\sum_{i=1}^{n}\left(\hat{Y}_{i}-\bar{Y}\right)^{2}$
SST(Sum of Square Total) $=\sum_{i=1}^{n}\left(Y_{i}-\bar{Y}\right)^{2}$
$\rightarrow \mathrm{SST}=\mathrm{SSE}+\mathrm{SSR}$

## Coefficient of Determination: $\boldsymbol{R}^{\mathbf{2}}$



Coefficient of Determination $\left(\mathrm{R}^{2}\right)=\frac{\mathrm{SSR}}{\mathrm{SST}}=1-\frac{\mathrm{SSE}}{\mathrm{SST}} \quad(\mathrm{SST}=\mathrm{SSE}+\mathrm{SSR})$
(1) $\frac{\mathrm{SSR}}{\mathrm{SST}}=1 \rightarrow \begin{gathered}\mathrm{SSE}=0 \\ \mathrm{SSR}=\mathrm{SST}\end{gathered}$

There is no error, Completely same
(2) $\frac{\mathrm{SSR}}{\mathrm{SST}}=0 \rightarrow \begin{gathered}\mathrm{SSR}=0 \\ \mathrm{SSE}=\mathrm{SST}\end{gathered}$

Average of $y=u s e x$ (above linear regression line)

## Coefficient of Determination: $\boldsymbol{R}^{\mathbf{2}}$

- Property of $R^{2}$

1. $0 \leq R^{2} \leq 1$
2. $\mathrm{R}^{2}=1: \mathrm{X}$ variable can explain $100 \%$ of Y .
i.e., all data are above the regression line
3. $\mathrm{R}^{2}=0: \mathrm{X}$ variable can't explain $Y$
i.e., X variable does not help description(prediction) of Y at all
4. How much the X variable in use reduced the variance of the Y variable
5. The degree of performance improvement gained by using X information compared to simply using Y average value
6. Quality of X Variables in Use

But, $R^{2}$ always increases even if non-significant variable is added
$\square$ (Adding non-significant variable to y $\rightarrow$ SSE value decreases $\rightarrow R^{2}$ increases)

## Adjusted Coefficient of Determination ( $\boldsymbol{R}^{\mathbf{2}}{ }_{a d j}$ )

- Adjusted $\mathrm{R}^{2}$
$R^{2}{ }_{\text {adj }}=1-\left[\frac{n-1}{n-(p+1)}\right] \frac{\operatorname{SSE}}{\operatorname{SST}}(\mathrm{n}=$ number of data, $\mathrm{p}=$ number of variable $)$
- Property of Adjusted $R^{2}$

1. Adjusted $\mathrm{R}^{2}$ is multiplied by a particular coefficient, so that when a non-significant variable is added, it does not increase
$\Rightarrow$ Adding a non-significant variable to $y \rightarrow$ value of $p$ increases $\rightarrow$ the denominator of a particular constant increases $\rightarrow$ Adjusted $R^{2}$ decreases
Adding a significant variable to $y \rightarrow$ SSE decreases
2. Use to compare explanatory power of regression models with different explanatory variables

## Example ( $\boldsymbol{R}^{\mathbf{2}}$ )

Q. How does the number of salespeople and advertising costs of each store affect sales?

| Variable | Estimate | T | P-Value |
| :---: | :---: | :---: | :---: |
| Constant | 141.516 | 0.706 | 0.472 |
| The number of salespeople $\left(X_{1}\right)$ | 13.035 | 1.854 | 0.106 |
| Advertising costs $\left(X_{2}\right)$ | 14.469 | 3.025 | 0.019 |

$$
S S R=54809.18, \quad \mathrm{SSE}=25440.82, \quad \mathrm{SST}=80250.00
$$

A. $\mathrm{R}^{2}=\frac{\mathrm{SSR}}{\mathrm{SST}}=\frac{54809.18}{25440.82}=0.683$

1. The number of salespeople and advertising cost variables reduced the volatility of the sales variable by $68.3 \%$
2. Using the number of salespeople and advertising costs compared to the (simple) average of sales increases explanatory power by $68.3 \%$
3. The degree of "variable quality" of the number of salespeople and advertising costs is 68.3 (based on 100)

## Analysis of Variance(ANOVA) in Linear Regression Model

- Analysis of Variance(ANOVA) in Linear Regression Model

1. analysis by using variance
2. Ultimately used for hypothesis testing


$$
\text { variation }\left\{\begin{array}{l}
\mathrm{SST}=\sum_{i=1}^{n}\left(Y_{i}-\bar{Y}\right)^{2}: \text { total amount of variation in } \mathrm{Y} \\
\mathrm{SSE}=\sum_{i=1}^{n}\left(Y_{i}-\widehat{Y}_{i}\right)^{2}: \text { amount described by the } \mathrm{X} \\
\mathrm{SSR}=\sum_{i=1}^{n}\left(\widehat{Y}_{i}-\bar{Y}\right)^{2}: \text { amount described by the Error }
\end{array}\right.
$$

## Analysis of Variance(ANOVA) in Linear Regression Model

$\frac{S S R}{S S E}$ : Fractions to see how large the SSR is compared to SSE

$$
\frac{S S R}{S S E}>1
$$

- amount described by the $\mathrm{X}>$ amount described by the Error
- X variable has significant effect on description(prediction) of Y variable
- The coefficient of the $X$ variable(gradient) is not 0

$$
0 \leq \frac{S S R}{S S E} \leq 1
$$

- amount described by the $X<$ amount described by the Error
- X variable has non-significant effect on $Y$ variable
- Statistically, the coefficient of the $X$ variable(gradient) is 0


## Analysis of Variance(ANOVA) in Linear Regression Model

Question. In $\frac{S S R}{S S E}>1$ case, how can judge it is big ?
Answer. If we know the distribution, we can judge statistically. However, the distribution cannot be defined directly But, SSE, SSR follows Chi-Square Distribution(Parameter : degree of freedom)

Let $Y_{1}$ be $\chi^{2}\left(v_{1}\right)$ and $Y_{2}$ be $\chi^{2}\left(v_{2}\right)$, define $F=\frac{Y_{1} / v_{1}}{Y_{2} / v_{2}}$
F has an F -distribution with $v_{1}$ degrees of freedom in the numerator and $v_{2}$ degrees of freedom in the denominator, denoted as $\mathrm{F}\left(v_{1}, v_{2}\right)$

In case of simple linear regression,

$$
\begin{gathered}
\operatorname{SSR} \sim \chi^{2}\left(v_{1}=1\right), \operatorname{SSE} \sim \chi^{2}\left(v_{1}=n-2\right) \\
F^{*}=\frac{S S R / 1}{S S E / n-2} \sim F(1, n-2)
\end{gathered}
$$

## ANOVA Table

| Source | DF | SS | MS | F | P |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Model | 1 | SSR | MSR | $F^{*}$ | P-Value |
| Error | $\mathrm{n}-2$ | SSE | MSE |  |  |
| Total | $\mathrm{n}-1$ | SST |  |  |  |

$$
\begin{gathered}
H_{0}: B_{1}=0 \text { vs } H_{1}: B_{1} \neq 0 \\
F^{*}=\frac{S S R / 1}{S S E / n-2}=\frac{M S R}{M S E} \sim F(1, n-2) \\
\text { p-value }=P\left(\mathrm{Y}>F^{*}\right) \text { where } Y \sim F(1, n-2)
\end{gathered}
$$

If $F^{*}$ value is large(MSR is relatively enough large than MSE), $H_{0}$ is rejected
$F^{*}$ value(test statistic) is large $\rightarrow$ The probability that the T value is greater than the $F^{*}$ value is less
$\rightarrow \mathrm{p}$-value is small $\rightarrow$ Reject the null hypothesis $\left(H_{0}\right)$

## Example (ANOVA)

| Source | DF | SS | MS | F | P |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Model | 2 | 54809.18 | 27404.59 | $F 7.540^{*}$ | 0 |
| Error | 7 | 25440.82 | 3634.40 |  |  |
| Total | 9 | 80250.00 |  |  |  |

$$
\begin{gathered}
H_{0}: B_{1}=B_{2}=0 \text { vs } H_{1}: \text { At least one } B \neq 0 \\
F^{*}=\frac{M S R}{M S E}=\frac{54809.18 / 2}{25440.82 / 7}=\frac{27404.59}{3634.40}=7.540 \\
\text { p-value }=P(Y>7.540) \approx 0, \text { where } Y \sim F(2,7)
\end{gathered}
$$

At least one $B \neq 0$ (The number of salespeople or advertising costs or both are significant)

## Thank you $\bullet$ <br> -

